### Recent Advances in Risk Aggregation and Dependence Uncertainty

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The Question	Mixability	Risk Aggregation	Challenges	References

Mainly based on some joint work with

- Valeria Bignozzi (Rome)
- Paul Embrechts (Zurich)
- Andreas Tsanakas (London)
- Bin Wang (Beijing)

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# Outline

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Risk aggregation				

Two aspects of modeling and inference of a multivariate model: **marginal distribution** and **dependence structure**.

"copula thinking"

- Assumption: certain margins, uncertain dependence.
- A common setup in operational risk

For example,

$$S_n = X_1 + \cdots + X_n.$$

 $X_i$ : individual risks;  $S_n$ : risk aggregation



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Simple example				

The simplest case: n = 2,  $F_1 = F_2 = U[-1, 1]$ . What is a possible distribution of  $S_2 = X_1 + X_2$ ?



**Obvious constraints** 

- $\mathbb{E}[S_2] = 0$
- range of  $S_2$  in [-2,2]
- $Var(S_2) \le 4/3$

• In fact,  $S_2 \prec_{cx} 2X_1$ i.e.  $S_2 \stackrel{d}{=} 2\mathbb{E}[X_1|\mathcal{G}]$ (sufficient?)

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This it not trivial any more<sup>1</sup>.

<sup>1</sup>the case [-1, 1] obtained in Rüschendorf (1982); general case [-a, a] obtained in Wang-W. (2015+ MOR)

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Example 3.3 of Mao-W. (2015 JMVA)	▲□▶ ▲圖▶ ▲圖▶ ▲圖▶	∃<∩<  <br< th=""></br<>

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Aggregation set				

Denote the aggregation set

$$\mathcal{D}_n = \mathcal{D}_n(F_1, \ldots, F_n) = \{ \text{cdf of } S_n | X_i \sim F_i, i = 1, \ldots, n \}.$$

•  $\mathcal{D}_n$  is a convex set, and closed with respect to weak convergence.

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Aggregation set				

Some questions to ask:

- (Compatibility) For a given F, is  $F \in \mathcal{D}_n$ ?
- (Mimicking) What is the best approximation in D<sub>n</sub> to F? That is, find G ∈ D<sub>n</sub> such that d(F, G) is minimized for some metric d.
- (Extreme values) What is  $\sup_{F \in D_n} \rho(F)$  for some functional  $\rho$ ?  $\leftarrow$  risk aggregation with dependence uncertainty

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Other application	IS			

Many applications and related problems

- Simulation: variance reduction
- Model-independent option pricing
- (Multi-dimensional) Monge-Kantorovich optimal transportation
- Change of measure
- Decision making
- Assembly and scheduling<sup>2</sup>

Many natural questions are not related to statistical uncertainty

<sup>2</sup>traditional problem in OR: e.g. Coffman-Yannakakis (1984 MOR) = ≥ + ( ≥ + ) ≥ → へへ Ruodu Wang (wang@uwaterloo.ca) Risk aggregation and dependence uncertainty 19/55

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Aggregation of ri	sk measures			

Attention coming from Quantitative Risk Management

 Most research looks at extreme values of some quantities (e.g. risk measures, pricing function) on the aggregate position S<sub>n</sub>:

$$\sup\{\rho(S_n): F_{S_n} \in \mathcal{D}\}$$
 and  $\inf\{\rho(S_n): F_{S_n} \in \mathcal{D}\}$ 

where  $\mathcal{D}$  is typically a subset of  $\mathcal{D}_n$ .

Earlier research:

VaR: Embrechts-Puccetti (2006 F&S)
 Distribution functions: Makarov (1981 TPA), Rüschendorf (1982 JAP)

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### Risk aggregation and dependence uncertainty

An active field for the past few years:

- Some recent papers (many more not listed)
  - W.-Peng-Yang (2013 F&S)
  - Embrechts-Puccetti-Rüschendorf (2013 JBF)
  - Bernard-Jiang-W. (2014 IME)
  - Aas-Puccetti (2014 Extremes)
  - Embrechts-Wang-W. (2015 F&S)
  - W.-Bignozzi-Tsanakas (2015 SIFIN)
  - Bignozzi-Puccetti-Rüschendorf (2015 IME)
  - Bernard-Vanduffel (2015 JBF)
  - Bernard-Vanduffel-Rüschendorf (2015+ JRI)
  - Wang-W. (2015+ MOR)

### Risk aggregation and dependence uncertainty

Books covering topics in this field:

### Rüschendorf (2013)



McNeil-Frey-Embrechts (2015)

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#### Observe that

$$S = X_1 + \cdots + X_n \Leftrightarrow X_1 + \cdots + X_n - S = 0$$

Hence,

$$F_{\mathcal{S}} \in \mathcal{D}_n(F_1,\ldots,F_n) \Leftrightarrow \delta_0 \in \mathcal{D}_{n+1}(F_1,\ldots,F_n,F_{-\mathcal{S}}).$$

To answer

is a distribution in  $\mathcal{D}_n$ ,  $n \geq 2$ ?

We study

is a point-mass in  $\mathcal{D}_{n+1}$ ,  $n \geq 2$ ?

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Joint mixability				

#### Joint mix

A random vector  $(X_1, \ldots, X_n)$  is a joint mix if  $X_1 + \cdots + X_n$  is a constant.

• Example: a multinomial random vector

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#### Joint mixability (W.-Peng-Yang, 2013 F&S)

An *n*-tuple of univariate distributions  $(F_1, \ldots, F_n)$  is jointly mixable (JM) if there exists a joint mix with marginal distributions  $(F_1, \ldots, F_n)$ .

- Equivalently,  $\mathcal{D}_n(F_1, \ldots, F_n)$  contains a point-mass.
- This concerned point-mass can be chosen at the sum of the means of F<sub>1</sub>,..., F<sub>n</sub> whenever it is finite.
- We say a univariate distribution *F* is *n*-completely mixabe (*n*-CM) if exists an *n*-dimensional joint mix with identical marginal distributions *F*.

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Mixability				

An open research area:

#### what distributions are CM/JM?

The research in this area is very much marginal-dependent - copula techniques do not help much!

• Recent summary paper: Puccetti-W. (2015 StS)

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Mean condition				

Let  $\mu_i, a_i, b_i \in \mathbb{R}$  be respectively the mean, essential infimum, and essential supremum of the support of  $F_i$ ;

$$I = \max\{b_i - a_i : i = 1, \ldots, n\}.$$

#### Mean condition

If  $(F_1, \ldots, F_n)$  is JM, then

$$\sum_{i=1}^{n} a_i + l \le \sum_{i=1}^{n} \mu_i \le \sum_{i=1}^{n} b_i - l$$
 (1)

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### Sufficiency of mean condition

### Sufficiency:

#### Theorem 1 (Wang-W., 2015+ MOR)

The mean condition (1) is sufficient for a tuple of distributions with increasing (decreasing) densities to be JM.

- The homogeneous case is shown in Wang-W. (2011 JMVA).
- Corollary:  $(\mathrm{U}[0,a_1],\ldots,\mathrm{U}[0,a_n])$  is JM if and only if

$$\max_{i=1,\ldots,n}a_i\leq \frac{1}{2}\sum_{i=1}^na_i.$$

• In particular<sup>3</sup>: U[0, 1] is *n*-CM for  $n \ge 2$ .

<sup>3</sup> known in Rüschendorf (1982 JAP)	▲□▶ ▲圖▶ ▲厘▶ ▲厘▶	æ	9 Q (?
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Variance conditi	ion			

#### Another necessary condition:

Variance condition

If  $(F_1, \ldots, F_n)$  is JM with finite variance  $\sigma_1^2, \ldots, \sigma_n^2$ , then

$$\max_{i=1,\dots,n} \sigma_i \le \frac{1}{2} \sum_{i=1}^n \sigma_i.$$
(2)

### (A polygon inequality<sup>4</sup>.)

<sup>4</sup>the standard deviation can be replaced by any law-based central norm  $\leftarrow$   $\equiv$   $\rightarrow$   $\bigcirc$   $\bigcirc$   $\bigcirc$   $\bigcirc$ 

# Sufficiency of variance condition

### Theorem 2 (Wang-W., 2015+ MOR)

The variance condition (2) is sufficient for the joint mixability of

- (i) a tuple of uniform distributions,
- (ii) a tuple of marginal distributions of a multivariate elliptical distribution,
- (iii) a tuple of distributions with unimodal-symmetric densities in the same location-scale family.

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Joint mixability				

#### Theorem 3 (Wang-W., 2015+ MOR)

Suppose that F has a unimodal-symmetric density. For a > 0, (U[0, a], U[0, a], F) is JM if and only if F is supported in an interval of length at most 2a.

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Joint mixability				

Some remarks:

- Determination of JM is still open
- 12 open questions on mixability: W. (2015 PS)
- Determination of JM in discrete setting is NP-complete<sup>5</sup>.

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### Risk Aggregation under Uncertainty

To study aggregation sets  $\mathcal{D}_n$ :

- To measure model uncertainty for quantities (e.g. risk measures, moments, etc) of *S*<sub>n</sub>.
- Targets:

$$\sup_{F_{S}\in\mathcal{D}_{n}}\rho(S) \text{ and } \inf_{F_{S}\in\mathcal{D}_{n}}\rho(S)$$
(3)

where  $\rho : \mathcal{X} \to \mathbb{R}$  is a risk measure<sup>6</sup>.

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ho$  is law-determined;  $\mathcal{X}$  is a set of random variables on  $(\Omega, \mathcal{F}_{\mathcal{F}})$   $\exists \mathcal{F} \in \mathbb{R}$ 

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VaR and ES				

#### Two regulatory risk measures

Value-at-Risk  $VaR_p$ 

For  $p \in (0,1)$ ,

$$\operatorname{VaR}_p(X) = F_X^{-1}(p) = \inf\{x \in \mathbb{R} : F_X(x) \ge p\}$$

#### Expected Shortfall $ES_p$

For  $p \in (0,1)$ ,

$$\mathrm{ES}_p(X) = \frac{1}{1-p} \int_p^1 \mathrm{VaR}_q(X) \mathrm{d}q \underset{(F \text{ cont.})}{=} \mathbb{E}\left[X | X > \mathrm{VaR}_p(X)\right]$$

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### Worst- and best-values of VaR and ES

The Fréchet (unconstrained) problems for  $\operatorname{VaR}_p$ : For given  $F_1, \ldots, F_n$  with finite means, and  $p \in (0, 1)$ , let

$$\overline{\mathrm{VaR}}_{p}(n) = \sup\{\mathrm{VaR}_{p}(S) : F_{S} \in \mathcal{D}_{n}(F_{1}, \ldots, F_{d})\},\$$

$$\underline{\operatorname{VaR}}_p(n) = \inf \{ \operatorname{VaR}_p(S) : F_S \in \mathcal{D}_n(F_1, \ldots, F_d) \}.$$

Same notation for  $ES_p$ .

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### Worst- and best-values of VaR and ES

Uncertainty intervals

 $[\underline{\operatorname{VaR}}_p(n), \overline{\operatorname{VaR}}_p(n)], \ [\underline{\operatorname{ES}}_p(n), \overline{\operatorname{ES}}_p(n)]$ 

- ES is subadditive:  $\overline{\mathrm{ES}}_p(n) = \sum_{i=1}^n \mathrm{ES}_p(X_i)$ .
- $\overline{\text{VaR}}_p(n)$ ,  $\underline{\text{VaR}}_p(n)$  and  $\underline{\text{ES}}_p(n)$ : generally open questions

Challenge for  $\underline{\mathrm{ES}}_{p}(n)$ 

To calculate  $\underline{\mathrm{ES}}_{p}(n)$  one naturally seeks a safest risk in  $\mathcal{D}_{n}$ .

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Mathematical dif	ficulty			

Common understanding of the most dangerous scenario:

 Comonotonicity - well accepted notion; even for a collection of random vectors

Understanding concerning the safest scenario:

- *n* = 2: counter-monotonicity
- $n \ge 3$ : unclear
  - Calls for notions of extremal negative dependence.

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Summary of existing results						

- n = 2: (based on counter-comonotonicity)
  - fully solved analytically<sup>7</sup>
- $n \ge 3$ : (based on joint mixability)
  - $\underline{\text{ES}}_{p}(n)$  solved analytically for decreasing densities, e.g. Pareto, Exponential
  - VaR<sub>p</sub>(n) solved analytically for tail-decreasing densities, e.g.
     Pareto, Gamma, Log-normal<sup>8</sup>
  - $\underline{\operatorname{VaR}}_p(n)$  similar to  $\overline{\operatorname{VaR}}_p(n)$

<sup>7</sup>Makarov (1981 TPA) and Rüschendorf (1982 JAP) <sup>8</sup>homogeneous model: W.-Yang-Peng (2013 F&S); inhomogeneous model: Jakobsons-Han-W. (2015+ SAJ)

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Remarks				

Remarks:

- For general marginal distributions the problem is still open
- Numerical methods: Rearrangement Algorithm<sup>9</sup>

<sup>9</sup>Puccetti-Rüschendorf (2012 JCAM); Embrechts-Puccetti-Rüschendorf (2013 JBF), Hofert-Memartoluie-Saunders-Wirjanto (2015+ arXiv)⇒ <♂⇒ < ≧⇒ < ≧⇒

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Aggregation of ris	sk measures			

Let  $\mathcal{D}_n(F) = \mathcal{D}_n(F, \dots, F)$  (homogeneous model). For a law-determined risk measure  $\rho$ , define

$$\Gamma_{\rho}(X) = \lim_{n \to \infty} \frac{1}{n} \sup \left\{ \rho(S) : F_S \in \mathcal{D}_n(F_X) \right\}.$$

 $\Gamma_{\rho}$  is also a law-determined risk measure which inherits some properties of  $\rho$ .

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Distortion risk measures:

$$\rho_h(X) = \int_0^1 F_X^{-1}(t) \mathrm{d}h(t), \ X \in \mathcal{X} = L^\infty$$

*h* is the distortion function: a probability measure on (0, 1).

• ES and VaR are special cases

Theorem 4 (W.-Bignozzi-Tsanakas, 2015 SIFIN)

We have

$$\Gamma_{\rho_h}(X) = \rho_{h^*}(X), \ X \in \mathcal{X},$$

where  $h^*$  is the largest convex distortion function dominated by h.

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### Aggregation of risk measures

For distortion risk measures

- Example:  $\Gamma_{\operatorname{VaR}_p} = \operatorname{ES}_p$
- $\rho_h$  is coherent if and only if  $h^* = h$
- Application: when arbitrary dependence is allowed, the worst-case VaR<sub>p</sub> of a portfolio behaves like the worst-case ES<sub>p</sub>

For law-determined convex risk measures.

- $\Gamma_{
  ho}$  is the smallest coherent risk measure dominating ho
- If  $\rho$  is a convex shortfall risk measure, then  $\Gamma_{\rho}$  is a coherent expectile

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### Dependence-uncertainty spread

Theorem 5 (Embrechts-Wang-W., 2015 F&S)

Take  $1 > q \ge p > 0$ . Under weak regularity conditions, for inhomogeneous models,

$$\liminf_{n\to\infty}\frac{\overline{\operatorname{VaR}}_q(n)-\underline{\operatorname{VaR}}_q(n)}{\overline{\operatorname{ES}}_p(n)-\underline{\operatorname{ES}}_p(n)}\geq 1.$$

- The uncertainty-spread of VaR is generally bigger than that of ES.
- In recent Consultative Documents of the Basel Committee, VaR<sub>0.99</sub> is compared with ES<sub>0.975</sub>: p = 0.975 and q = 0.99.

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### Dependence-uncertainty spread

ES and VaR of 
$$S_n = X_1 + \cdots + X_n$$
, where

• 
$$X_i \sim \text{Pareto}(2 + 0.1i), \ i = 1, \dots, 5;$$

• 
$$X_i \sim \text{Exp}(i-5), i = 6, ..., 10;$$

• 
$$X_i \sim \text{Log-Normal}(0, (0.1(i-10))^2), i = 11, \dots, 20.$$

		<i>n</i> = 5			<i>n</i> = 20	
	best	worst	spread	best	worst	spread
ES <sub>0.975</sub>	22.48	44.88	22.40	29.15	102.35	73.20
$VaR_{0.975}$	9.79	41.46	31.67	21.44	100.65	79.21
$VaR_{0.99}$	12.96	62.01	49.05	22.29	136.30	114.01
$\frac{\overline{\mathrm{ES}}_{0.975}}{\overline{\mathrm{VaR}}_{0.975}}$		1.08			1.02	

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### Dependence-uncertainty spread

Features/Risk measure	VaR	Tail-VaR
Frequency captured?	Yes	Yes
Severity captured?	No	Yes
Sub-additive?	Not always	Always
Diversification captured?	Issues	Yes
Back-testing?	Straight-forward	Issues
Estimation?	Feasible	Issues with data limitation
Model uncertainty?	Sensitive to aggregation	Sensitive to tail modelling
Robustness I (with respect to "Lévy metric <sup>33</sup> ")?	Almost, only minor issues	No
Robustness II (with respect to "Wasserstein metric <sup>34</sup> ")?	Yes	Yes

From the International Association of Insurance Supervisors Consultation Document (December 2014).

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Open questions				
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Concerete mathematical questions:

- Full characterization of  $\mathcal{D}_n$  and mixability
- Existence and determination of smallest  $\prec_{\mathrm{cx}}$ -element in  $\mathcal{D}_n$
- General analytical formulas for  $\overline{\mathrm{VaR}}_p$  ( $\underline{\mathrm{VaR}}_p$ ) and  $\underline{\mathrm{ES}}_p$
- Aggregation of random vectors

Practical questions:

- Capital calculation under uncertainty
- Robust decision making under uncertainty
- Regulation with uncertainty

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Other directions				

Some on-going directions on RADU

- Partial information on dependence<sup>10</sup>
- Connection with Extreme Value Theory
- Connection with martingale optimal transportation
- Both marginal and dependence uncertainty
- Computational solutions
- Other aggregation functionals

<sup>10</sup>Bignozzi-Puccetti-Rüschendorf (2015 IME), Bernard-Rüschendorf-Vanduffel (2015+ JRI), Bernard-Denuit-Vanduffel (2014 SSRN), Bernard-Vanduffel (2015 JBF), many more

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## Thank you for your kind attention